

Marginalization in Lazy Propagation

Alg. 3.1.1 [Madsen's dissertation]

Let $\phi = \{\varphi_1, \dots, \varphi_n\}$ be a set of potentials.

If marginalization of X is invoked on ϕ , then

1. set $\phi_x = \{\varphi \in \phi \mid X \in \text{dom}(\varphi)\}$

2. $\varphi_x^* = \sum_{\varphi \in \phi_x} \pi \varphi$

3. $\phi^* = \{\varphi_x^*\} \cup \phi \setminus \phi_x$

ϕ^* is the set of potentials that results from

eliminating X from ϕ . □

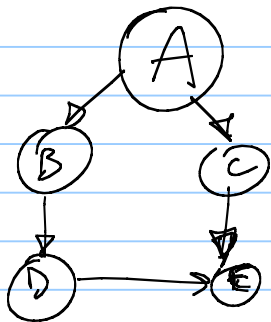
This is Defn. 4.1 on p. 166 with different terminology.

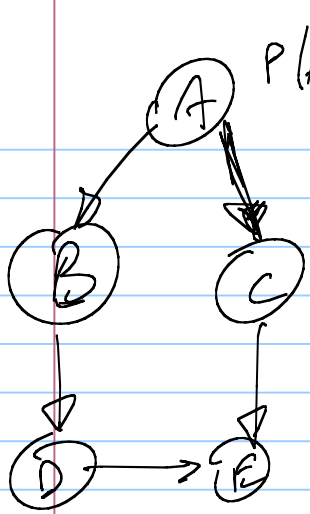
elimination for marginalization

ϕ^{-x} for x^* .

4.8 Stochastic Simulation in Bayesian Networks

$$P(E=y) \approx \frac{N(E=y)}{N} = \frac{\text{number of cases in which } E=y}{\text{total number of cases}}$$





$$P(A) = (.4, .6)$$

	A	n
B	.4	.8
n	.3	.2

C	A	n
y	.7	.4
n	.3	.6

D	B	n
y	.5	.1
n	.5	.9

	C	n
D	.4	
y	(.9, .1)	(.999, .001)
n	(.999, .001)	(.999, .001)

$\rightarrow P(E|C, D)$

CDR

AB	yy	yn	ny	nn
yy	4	0	5	0
yn	2	0	16	0
ny	9	1	10	0
nn	0	0	4	0

$$P(B=y) \approx \frac{N(B=y)}{N} = \frac{62}{100} = .62$$

$$P(E) \approx \left(\frac{N(E=y)}{N}, \frac{N(E=n)}{N} \right) = \left(\frac{99}{100}, \frac{1}{100} \right) = .99, .01$$

The algorithm (p. 48 [JOZ]).

1. Let $\langle x_1, \dots, x_n \rangle$ be a topological ordering of the variables (e.g. $\langle A, B, C, D, E \rangle$)
2. For $j=1$ to N :
 - a) For $i=1$ to n :
 - sample a state x_i of X_i using $P(X_i | \text{pe}(X_i) = \pi)$, where π is the configuration already sampled for $\text{pe}(X_i)$.
 - b) If $\underline{x} = (x_1, \dots, x_n)$ is consistent with \underline{e} , then

$N(X_k = x_k) := N(X_k = x_k) + 1$, where
 x_k is the state that was sampled for X_k

[else: do score $x \rightarrow$ because it is inconsistent with
the evidence]

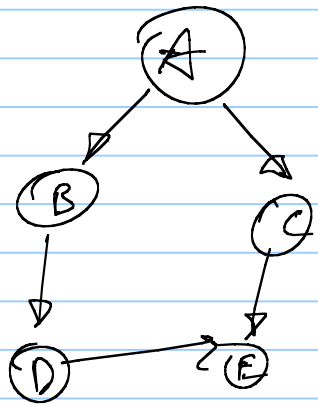
3. Return
$$P(X_k = x_k | e) = \frac{N(X_k = x_k)}{\sum_{x \in \text{sp}(X_k)} N(X_k = x)}$$

$$P(X_k = x_k | e) = \frac{P(X_k = x_k, e)}{P(e)}$$

The alg. described requires generating a lot of
irrelevant samples when the evidence has low probability.

4.8.2 Likelihood weighting was designed to overcome this problem.

Idea: to avoid flipping a biased coin for each variable on which there is evidence. *Wrong results!*



$$P(A | B=n, E=n)$$

By following the idea, you estimate

$P(A)$, not $P(A | \underline{e})$, b/c when the state of A is determined, only $P(A)$ is used

The idea can be salvaged by carefully writing down

$$(4.4) P(\underline{u}, \underline{e}) = \prod_{X \in U \setminus E} P(X | pa(X)', pa(X)'' = e) \times \prod_{X \in E} P(X=e | pa(X)', pa(X)'' = e)$$

E is the set of variables that have received evidence

The mistaken (naive) algorithm (based on the "blee") ignores the second part!

So, the fix is to weigh each sample by

$$\prod_{X \in \mathcal{E}} P(X=e | p_e(X), p_e(X) = e)$$

$$w(\underline{x}, \underline{e}) = \prod_{E \in \mathcal{E}} P(E=e | p_e(E) = \pi)$$

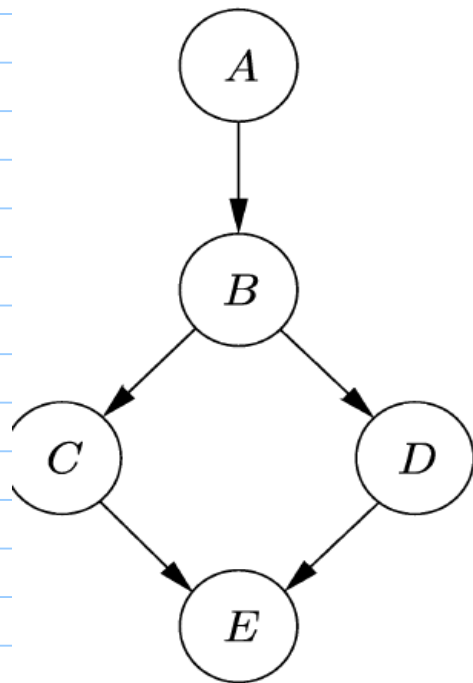
configuration in the sample point

In practice, once a sample point (case) has been produced, $w(\underline{x}, \underline{e})$ is calculated and it is added to $N(?)$ instead of 1.

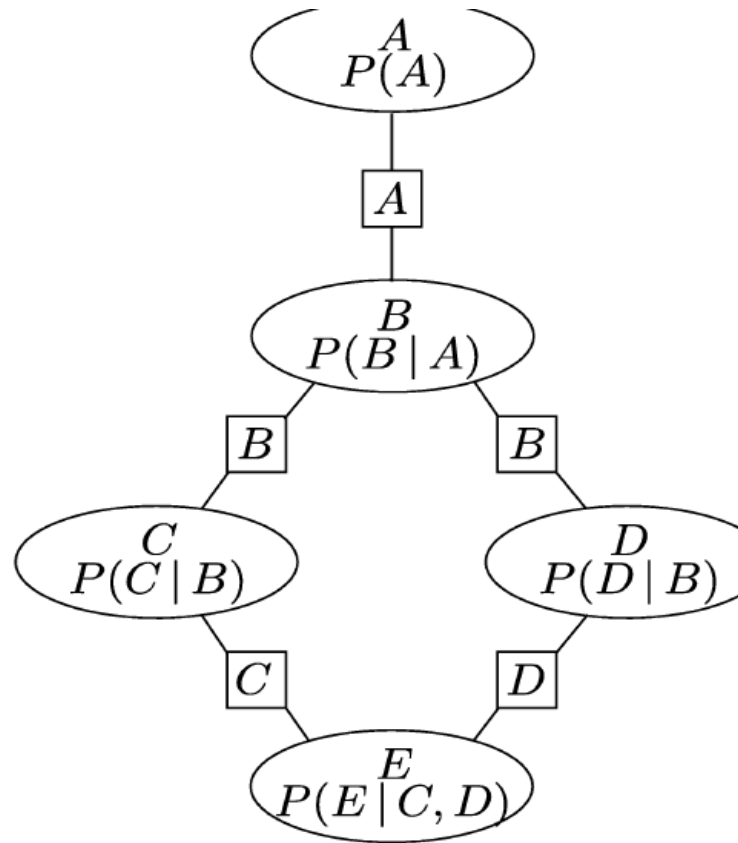
4.8.3 Gibbs Sampling (used Nielsen slides, unpublished)

4.9 Loopy Belief Propagation (~1998?)

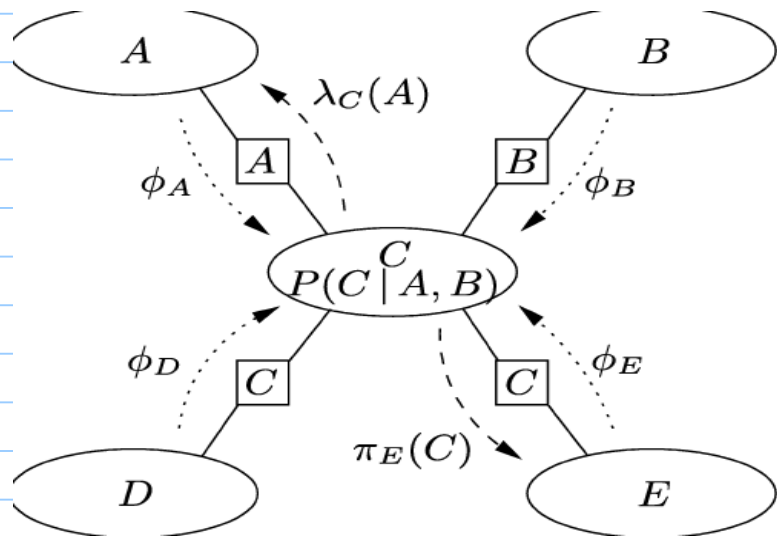
Used in "Turbo
Coding"
(McEliece)



(a)



(b)



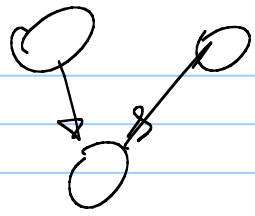
$$\lambda_C(A) = \sum_{B,C} P(C|A, B) \phi_B \phi_D \phi_E$$

$$\pi_E(C) = \phi_D \sum_{A,B} P(C|A, B) \phi_A \phi_B$$

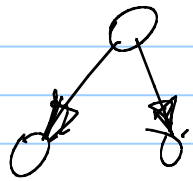
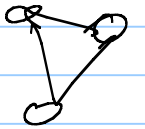
In the special case of BNs that are trees, this algorithm is the same as the junction tree algorithm.

The alg. in the case of trees is due to Pearl (& Shachter) (1982)

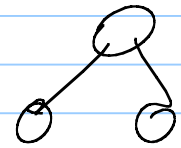
The alg also works for polytrees - BNs whose moral graph is a singly connected graph.



not a polytree, b/c its moral graph is



a polytree, b/c its moral graph is



By the observation that each clique in a j-tree for a polytree is small, you may show that belief updates in polytrees can be done in linear time.